

$$2^{\text{nd}} \text{ NED} : k : P(2 \text{nd op } H_0 | \theta = \theta_0) = \alpha$$

$$2^{\text{nd}} \text{ NED} : k : k_1 P(2 \text{nd op } H_0 | H_0 \text{ \& } \omega) = k_0 \cdot P(2 \text{nd op } H_0 | H_0 \text{ \& } \omega)$$

$$3^{\text{rd}} \text{ NED} : k = \frac{2(\theta_0, d_0) \cdot \pi(\theta_0)}{2(\theta_0, d_0) \cdot \pi(\theta_0)} = \frac{k_0}{k_1} \cdot \frac{\pi(\theta_0)}{\pi(\theta_0)}$$

Παράδειγμα 1

$$X_1, X_2, \dots, X_n \sim N(\theta, 100)$$

$$n = 100$$

$$H_0 : \theta = 75 \vee H_1 : \theta = 78$$

i) Ισχυρότητα τεστ 5%.

$$\text{ii) MinMax } 2(\theta_0, d_1) = 3 \text{ και}$$

$$\text{MinMax } 2(\theta_1, d_0) = 1$$

$$\text{iii) Bayes τεστ } 2(\theta_0, d_1) = 2(\theta_1, d_0)$$

$$\pi(\theta_0) = \pi(\theta_1)$$

Να βρεθεί το α και το γ .

Απάντηση

$$\text{i) } \frac{\pi(x_i | \theta_1)}{\pi(x_i | \theta_0)} \leq k \Rightarrow \frac{e^{-\frac{1}{2\sigma^2} \sum (x_i - \theta_1)^2}}{e^{-\frac{1}{2\sigma^2} \sum (x_i - \theta_0)^2}} \leq k \Rightarrow$$

$$\Rightarrow e^{-\frac{1}{2\sigma^2} \{-2\theta_0 \sum x_i + n\theta_0^2 - 2\theta_1 \sum x_i + n\theta_1^2\}} \leq k \quad (1)$$

$$\Rightarrow -\frac{1}{2\sigma^2} (2 \sum x_i (\theta_1 - \theta_0) + n(\theta_0^2 - \theta_1^2)) \leq \log k \Rightarrow$$

$$\begin{aligned}
 2 \sum x_i (\theta_1 - \theta_0) + n (\theta_0^2 - \theta_1^2) &\geq -2\sigma^2 \log k \Rightarrow \\
 \Rightarrow \sum x_i (\theta_1 - \theta_0) &\geq -\sigma^2 \log k - \frac{n}{2} (\theta_0^2 - \theta_1^2) \Rightarrow \\
 \Rightarrow \sum x_i &\geq \frac{-\sigma^2 \log k - \frac{n}{2} (\theta_0^2 - \theta_1^2)}{\theta_1 - \theta_0} = k_L
 \end{aligned}$$

k : τιω n πιθανότητα σφ. ενου I

$$\begin{aligned}
 P(\text{απορ } H_0 / H_0 \text{ αληθ}) &= \alpha \Rightarrow P(\sum x_i \geq k_L / X_i \sim N(\theta_0, \sigma^2) \text{ } | x_i) = \alpha \\
 \Rightarrow P(\sum x_i \geq k_L / \sum x_i \sim N(n\theta_0, n\sigma^2)) &= \alpha \Rightarrow \\
 \Rightarrow P(\frac{\sum x_i - n\theta_0}{\sqrt{n\sigma^2}} \geq \frac{k_L - n\theta_0}{\sqrt{n\sigma^2}} / H_0 \text{ αληθ}) &= \alpha
 \end{aligned}$$

$$\text{οπου } Z_\alpha = \frac{k_L - n\theta_0}{\sqrt{n\sigma^2}} \Rightarrow k_L = \sqrt{n\sigma^2} Z_\alpha + n\theta_0$$

για την ωχύ του τεστ:

$$\begin{aligned}
 \gamma &= P(\text{απορ } H_0 / H_a \text{ αληθ}) = P(\sum x_i \geq k_L / \sum x_i \sim N(n\theta_a, n\sigma^2)) = \\
 &= P(\frac{\sum x_i - n\theta_a}{\sqrt{n\sigma^2}} \geq \frac{k_L - n\theta_a}{\sqrt{n\sigma^2}} / \sum x_i \sim N(n\theta_a, n\sigma^2)) = \\
 &= 1 - \Phi\left(\frac{k_L - n\theta_a}{\sqrt{n\sigma^2}}\right).
 \end{aligned}$$

$$\text{ii) } k_L \cdot P(\sum x_i \geq k_L / X_i \sim N(\theta_0, \sigma^2)) = k_0 \cdot P(\sum x_i < k_L / X_i \sim N(\theta_a, \sigma^2))$$

$$\begin{aligned}
 3. P(\frac{\sum x_i - n\theta_0}{\sqrt{\sigma^2 \cdot n}} \geq \frac{k_L - n\theta_0}{\sqrt{\sigma^2 \cdot n}} / H_0 \text{ αληθ}) &= \\
 = 1 \cdot P(\frac{\sum x_i - n\theta_a}{\sqrt{\sigma^2 \cdot n}} < \frac{k_L - n\theta_a}{\sqrt{\sigma^2 \cdot n}} / H_a \text{ αληθ}) &\Rightarrow \\
 \Rightarrow 3 \cdot (1 - \Phi(\frac{k_L - n\theta_0}{\sqrt{\sigma^2 \cdot n}})) &= 1 \cdot \Phi(\frac{k_L - n\theta_a}{\sqrt{\sigma^2 \cdot n}}) \quad (*)
 \end{aligned}$$

Αρκεί, να βρω το k_L για το οποίο επαληθεύεται η εξίσωση (*). Η σωθίκεν (*) θα δίνεται στις εξετάσεις σε συνένδετα ένα συκτεμπλεω k_L . Στην προκειμένη περίπτωση το $k_L = 7680$.

Απάντηση στο ερώτημα: Ποιο το επίπεδο συκταμικότητας;

$$\alpha = P(\sum X_i \geq k_1 \mid X_i \sim N(\theta_0, \sigma^2))$$

και δειξαμε πριν, ου τελικα:

$$\alpha = 1 - \Phi\left(\frac{k_1 - n\theta_0}{\sqrt{n}\sigma}\right) \text{ για } k_1 \text{ δοδεν.}$$

Για των ωχου

$$\gamma = 1 - \beta = 1 - \Phi\left(\frac{k_1 - n\theta_1}{\sqrt{n}\sigma}\right) \text{ για } k_1 \text{ δοδεν}$$

$$\text{iii) Εάν } \lambda = e^{-\frac{1}{2\sigma^2} \{\sum (x_i - \theta_0)^2 - \sum (x_i - \theta_a)^2\}} \leq k = \frac{k_0}{k_1} \cdot \frac{\pi(\theta_a)}{\pi(\theta_0)}$$

$$\Rightarrow \lambda \leq \frac{k_0}{k_1} \cdot \frac{\pi(\theta_a)}{\pi(\theta_0)} = 1 \Rightarrow -\frac{1}{2\sigma^2} \{\sum (x_i - \theta_0)^2 - \sum (x_i - \theta_a)^2\} \leq 0$$

$$\Rightarrow \sum (x_i - \theta_0)^2 - \sum (x_i - \theta_a)^2 \geq 0 \Rightarrow$$

$$\Rightarrow -2 \sum x_i \theta_0 + n \theta_0^2 + 2 \sum x_i \theta_a - n \theta_a^2 \geq 0 \Rightarrow$$

$$\Rightarrow 2 \sum x_i (\theta_a - \theta_0) \geq n \theta_a^2 - n \theta_0^2 \Rightarrow$$

$$\Rightarrow \sum x_i \geq \frac{n(\theta_a^2 - \theta_0^2)}{2(\theta_a - \theta_0)} = \frac{n(\theta_a + \theta_0)}{2} = \frac{100(78+75)}{2} = 7650$$

Οπως, και πριν

$$\alpha = P(\text{απορ } H_0 \mid H_0 \text{ αληθης}) = \dots = 1 - \Phi\left(\frac{k_2 - n\theta_0}{\sqrt{n}\sigma}\right)$$

και

$$\gamma = P(\text{απορ } H_0 \mid H_1 \text{ αληθης}) = \dots = 1 - P(\sum x_i \leq k_2 \mid \sum x_i \sim N(n\theta_1, n\sigma^2))$$

Παράδειγμα 2

$X_1, X_2, \dots, X_n$ εδ απο των $\text{Exp}(\theta)$

Να ελεγχετε τις υποθεσεις

$$H_0: \theta = \theta_0 = 1 \quad \vee \quad H_1: \theta = \theta_1 = 2$$

i) ωχυροτατο τεστ $\alpha = 5\%$

ii) Min max: $\lambda(\theta_0, d_1) = \lambda(\theta_1, d_0)$

$$\text{iii) } \frac{\pi(\theta_1)}{\pi(\theta_2)} = \frac{1}{2}$$

$$\frac{\lambda_{\theta_0}}{\lambda_{\theta_1}} = \frac{\prod_{i=1}^n \theta_0 e^{-\theta_0 x_i}}{\prod_{i=1}^n \theta_1 e^{-\theta_1 x_i}} = \frac{\theta_0^n}{\theta_1^n} \frac{e^{-\theta_0 \sum_{i=1}^n x_i}}{e^{-\theta_1 \sum_{i=1}^n x_i}} = \frac{\theta_0^n}{\theta_1^n} e^{-\sum_{i=1}^n x_i (\theta_0 - \theta_1)}$$

$$i) \lambda \leq k \Rightarrow \frac{\theta_0^n}{\theta_1^n} e^{-\sum_{i=1}^n x_i (\theta_0 - \theta_1)} \leq k \Rightarrow$$

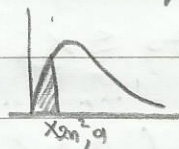
$$\Rightarrow e^{-\sum_{i=1}^n x_i (\theta_0 - \theta_1)} \leq \frac{k \theta_1^n}{\theta_0^n} \Rightarrow$$

$$\Rightarrow -\sum_{i=1}^n x_i (\theta_0 - \theta_1) \leq \log \left(\frac{k \theta_1^n}{\theta_0^n} \right) \Rightarrow$$

$$\Rightarrow \sum_{i=1}^n x_i \leq \frac{\log \left(\frac{k \theta_1^n}{\theta_0^n} \right)}{\theta_1 - \theta_0} = k_1 \quad (*)$$

η ισοδύναμη $\boxed{\sum x_i \leq k_1}$

$$\begin{aligned} i) \alpha &= P(\sum x_i \leq k_1 \mid x_i \sim \text{Eκθ}(\theta_0), \forall i) = \\ &= P(\sum x_i \leq k_1 \mid x_i \sim \text{Γαλλα}(1, \frac{1}{\theta_0})) = \\ &= P(\sum x_i \leq k_1 \mid x_i \sim \text{Γαλλα}(n, \frac{1}{\theta_0})) = \\ &= P(2\theta_0 \sum x_i \leq 2\theta_0 k_1 \mid 2\theta_0 x_i \sim \chi_{2n}^2) \Rightarrow \\ &\Rightarrow 2\theta_0 k_1 = \chi_{2n, \alpha}^2 \end{aligned}$$



Για την ποσότητα του ερεσι:

$$\gamma = P(\sum x_i \leq \frac{\chi_{2n, \alpha}^2}{2\theta_0} \mid x_i \sim \text{Eκθ}(\theta_1)) =$$

$$= P(\sum x_i \leq \frac{\chi_{2n, \alpha}^2}{2\theta_0} \mid \sum x_i \sim \text{Γαλλα}(n, \frac{1}{\theta_1})) =$$

$$= P(2\theta_1 \sum x_i \leq \frac{\chi_{2n, \alpha}^2}{2\theta_0} \cdot 2\theta_1 \mid 2\theta_1 \sum x_i \sim \chi_{2n}^2) =$$

$$= F_{\chi_{2n}^2} \left(\frac{\theta_1}{\theta_0} \cdot \chi_{2n, \alpha}^2 \right). \quad (\text{το οποίο θα δίνεται στις εξετάσεις})$$

ii) Το k_1 προσδιορίζεται έτσι ώστε

$$\begin{aligned}
 d(\theta_0, d_1) \cdot P(\text{απορ. } H_0 \mid H_0 \text{ αληθ.}) &= d(\theta_1, d_0) \cdot P(\text{απορ. } H_0 \mid H_0 \text{ αληθ.}) \\
 P(\sum X_i \leq k_1 \mid X_i \sim \text{Gamma}(1, \frac{1}{\theta_0})) &= P(\sum X_i > k_1 \mid X_i \sim \text{Gamma}(1, \frac{1}{\theta_1})) \\
 P(\sum X_i \leq k_1 \mid \sum X_i \sim \text{Gamma}(n, \frac{1}{\theta_0})) &= P(\sum X_i > k_1 \mid \sum X_i \sim \text{Gamma}(n, \frac{1}{\theta_1})) \\
 P(2\theta_0 \sum X_i \leq 2\theta_0 k_1 \mid 2\theta_0 \sum X_i \sim \chi_{2n}^2) &= P(2\theta_1 \sum X_i > 2\theta_1 k_1 \mid 2\theta_1 \sum X_i \sim \chi_{2n}^2) \\
 F_{\chi_{2n}^2}(2\theta_0 k_1) &= 1 - F_{\chi_{2n}^2}(2\theta_1 k_1) \quad (\text{αν μας δίνεται σε ποιο} \\
 &\quad \text{κ. δίνεται η συνάρτηση αυτή, αν όχι τότε το αφαιρούμε} \\
 &\quad \text{σε αυτή τη μορφή})
 \end{aligned}$$

$$\text{iii) } k = \frac{1}{2}, \quad (\text{αφού } d(\theta_0, d_0) = d(\theta_1, d_0) \text{ και } \frac{\pi(\theta_1)}{\pi(\theta_0)} = \frac{1}{2})$$

Αρα, από την (*)

$$\sum X_i \leq \frac{\log\left(\frac{1}{2} \frac{2^n}{1^n}\right)}{2-1} \Rightarrow \sum X_i \leq \log 2^n \Rightarrow$$

$$\Rightarrow \sum X_i \leq (n-1) \log 2 \quad \text{η καλύτερη περίπτωση}$$

$$\begin{aligned}
 \alpha &= P(\sum X_i \leq (n-1) \log 2 \mid X_i \sim \text{Gamma}(1, \frac{1}{\theta_0})) = \\
 &= P(2\theta_0 \leq X_i \leq 2\theta_0(n-1) \log 2) = \\
 &= F_{\chi_{2n}^2}(2\theta_0(n-1) \log 2)
 \end{aligned}$$

Λογός

$$\begin{aligned}
 \gamma &= P(\sum X_i \leq (n-1) \log 2 \mid X_i \sim \text{Gamma}(1, \frac{1}{\theta_1})) = \\
 &= P(2\theta_1 \sum X_i \leq 2\theta_1(n-1) \log 2)
 \end{aligned}$$

Δοκίμια

$X_1, X_2, \dots, X_n$ τ.δ. με $n=10$ από Poisson (θ)

Να ελεγχουμε $H_0: \theta = \theta_0 = 0.5$ vs $H_1: \theta = \theta_1 = 1$,

και να βρεθεί:

i) βέλτιστο τεστ με $3\% < \alpha < 5\%$.

ii) Minimax τεστ: $d(\theta_0, d_1) = 1.65$ και $d(\theta_1, d_0) = 1$

iii) Bayes τεστ: $d(\theta_1, d_1) = 2$ και $d(\theta_0, d_1) = 1$

$$\pi(\theta_0) = \frac{1}{3} \text{ και } \pi(\theta_1) = \frac{2}{3}$$

Λύση

$$\lambda := \frac{\prod_{i=1}^n \frac{\theta_0^{x_i} \cdot e^{-\theta_0}}{x_i!}}{\prod_{i=1}^n \frac{\theta_1^{x_i} \cdot e^{-\theta_1}}{x_i!}} = \left(\frac{\theta_0}{\theta_1} \right)^{\sum x_i} \cdot e^{-n(\theta_0 - \theta_1)} \leq k$$

$$\left(\frac{\theta_0}{\theta_1} \right)^{\sum x_i} \leq k \cdot e^{n(\theta_0 - \theta_1)} \Rightarrow \sum x_i \log \left(\frac{\theta_0}{\theta_1} \right) \leq \log k + n(\theta_0 - \theta_1)$$

$$\Rightarrow \sum x_i \geq \frac{\log(k \cdot e^{n(\theta_0 - \theta_1)})}{\log \frac{\theta_0}{\theta_1}} = k^* \quad (*)$$

i) βέλτιστο τεστ:

$$\alpha = P(\sum x_i \geq k^* \mid X_i \sim \text{Poisson}(\theta_0 = 0.5)) =$$

$$= P(\sum x_i \geq k^* \mid \sum x_i \sim \text{Poisson}(n \cdot \theta_0 = 5)) =$$

$$= 1 - P(\sum x_i < k^* \mid \sum x_i \sim \text{Poisson}(5)) =$$

$$= 1 - P(\sum x_i < k^* - 1 \mid \sum x_i \sim \text{Poisson}(5))$$

Επών $\alpha \in (3\%, 5\%)$ από τον πίνακα της Poisson κατανομής ένα τέτοιο α συν 5.0 είναι το 0.9682 . Επομένως, $k^* - 1 = 9 \Rightarrow k^* = 10$.

Αρα, $\sum X_i \geq 10$ η κρίση είναι

H ισχύει:

$$\gamma = P(\sum X_i \geq 10 \mid \sum X_i \sim \text{Poisson}(n\theta = 10)) =$$

$$= 1 - P(\sum X_i < 10 \mid \sum X_i \sim \text{Poisson}(10)) =$$

$$= 1 - P(\sum X_i \leq 9 \mid \sum X_i \sim \text{Poisson}(10)) =$$

$$= 1 - 0,4579.$$

$$\text{iii)} \quad k = \frac{2}{1} \cdot \frac{2/3}{1/3} = 4$$

Αρα η σχέση (*) δίνει

$$\sum X_i \geq \frac{\log(4 \cdot e^{10 \cdot (0,5-1)})}{\log \frac{0,5}{1}} = \frac{\log(4 \cdot e^{-5})}{-\log 2} =$$

$$= \frac{2\log 2 - 5}{-\log 2} = -2 + \frac{5}{\log 2} = 5,2$$

$$\text{Αρα, } \sum X_i \geq 5,2 \Leftrightarrow \sum X_i \geq 6.$$

$$\alpha = P(\sum X_i \geq 6 \mid \sum X_i \sim \text{Poisson}(5)) =$$

$$= 1 - P(\sum X_i < 6 \mid \sum X_i \sim \text{Poisson}(5)) =$$

$$= 1 - P(\sum X_i \leq 5 \mid \sum X_i \sim \text{Poisson}(5)) =$$

$$= 1 - 0,6160$$

H ισχύει:

$$\gamma = P(\sum X_i \geq 6 \mid \sum X_i \sim \text{Poisson}(10)) =$$

$$= 1 - P(\sum X_i \leq 5 \mid \sum X_i \sim \text{Poisson}(10)) =$$

$$= 1 - 0,0671$$

$$ii) 1,65 \cdot P(\sum X_i \geq k^* / X_i \sim \text{Poisson}(0,5)) = 1 \cdot P(\sum X_i < k^* / X_i \sim P(1))$$

$$1,65 P(\sum X_i \geq k^* / \sum X_i \sim \text{Poisson}(5)) = 1 \cdot P(\sum X_i < k^* / \sum X_i \sim P(10))$$

$$1,65 (1 - P(\sum X_i \leq k^* - 1 / \sum X_i \sim P(5))) = P(\sum X_i \leq k^* - 1 / \sum X_i \sim P(10))$$

Η λύση θα βρεθεί για εκείνο το k^* που των επαληθεύει
Μέσω δοκιμών:

$$\bullet k^* - 1 = 0 \Rightarrow k^* = 1$$

$$\bullet k^* - 1 = 1 \Rightarrow k^* = 2$$

$$\bullet k^* - 1 = 2 \Rightarrow k^* = 3$$

⋮

$$\bullet k^* - 1 = 8 \Rightarrow k^* = 9$$

⋮

Για $k^* - 1 = 0 \Rightarrow k^* = 1$ είναι:

$$1,65 \cdot (1 - 0,0067) = 0$$

$$1,65 \cdot (1 - 0,0404) = 0,005$$

$$1,65 \cdot (1 - 0,1247) = 0,0028$$

⋮

Μετά από επαληθεύσεις θα βρούμε ότι $k^* = 8$.

Αυτή που θα ικανοποιεί των εξισώσεων παραπάνω

θα είναι και μοναδική

Tables of the Poisson Cumulative Distribution

The table below gives the probability of that a Poisson random variable X with mean $= \lambda$ is less than or equal to x . That is, the table gives

$$P(X \leq x) = \sum_{r=0}^x \lambda^r \frac{e^{-\lambda}}{r!}$$

λ	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	1.0	1.2	1.4	1.6	1.8
0	0.9048	0.8187	0.7408	0.6703	0.6065	0.5488	0.4966	0.4493	0.4066	0.3679	0.3012	0.2466	0.2019	0.1653
1	0.9953	0.9825	0.9631	0.9384	0.9098	0.8781	0.8442	0.8088	0.7725	0.7358	0.6626	0.5918	0.5249	0.4628
2	0.9998	0.9989	0.9964	0.9921	0.9856	0.9769	0.9659	0.9526	0.9371	0.9197	0.8795	0.8335	0.7834	0.7306
3	1.0000	0.9999	0.9997	0.9992	0.9982	0.9966	0.9942	0.9909	0.9865	0.9810	0.9662	0.9463	0.9212	0.8913
4	1.0000	1.0000	1.0000	0.9999	0.9998	0.9996	0.9992	0.9986	0.9977	0.9963	0.9923	0.9857	0.9763	0.9636
5	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9997	0.9994	0.9985	0.9968	0.9940	0.9896
6	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9999	0.9997	0.9994	0.9987	0.9974
7	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997	0.9994
8	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
9	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
λ	2.0	2.2	2.4	2.6	2.8	3.0	3.2	3.4	3.6	3.8	4.0	4.5	5.0	5.5
0	0.1353	0.1108	0.0907	0.0743	0.0608	0.0498	0.0408	0.0334	0.0273	0.0224	0.0183	0.0111	0.0067	0.0041
1	0.4060	0.3546	0.3084	0.2674	0.2311	0.1991	0.1712	0.1468	0.1257	0.1074	0.0916	0.0611	0.0404	0.0266
2	0.6767	0.6227	0.5697	0.5184	0.4695	0.4232	0.3799	0.3397	0.3027	0.2689	0.2381	0.1736	0.1247	0.0884
3	0.8571	0.8194	0.7787	0.7360	0.6919	0.6472	0.6025	0.5584	0.5152	0.4735	0.4335	0.3423	0.2650	0.2017
4	0.9473	0.9275	0.9041	0.8774	0.8477	0.8153	0.7806	0.7442	0.7064	0.6678	0.6288	0.5321	0.4405	0.3575
5	0.9834	0.9751	0.9643	0.9510	0.9349	0.9161	0.8946	0.8705	0.8441	0.8156	0.7851	0.7029	0.6160	0.5289
6	0.9955	0.9925	0.9884	0.9828	0.9756	0.9665	0.9554	0.9421	0.9267	0.9091	0.8893	0.8311	0.7622	0.6860
7	0.9989	0.9980	0.9967	0.9947	0.9919	0.9881	0.9832	0.9769	0.9692	0.9599	0.9489	0.9134	0.8666	0.8095
8	0.9998	0.9995	0.9991	0.9985	0.9976	0.9962	0.9943	0.9917	0.9883	0.9840	0.9786	0.9597	0.9319	0.8944
9	1.0000	0.9999	0.9998	0.9996	0.9993	0.9989	0.9982	0.9973	0.9960	0.9942	0.9919	0.9829	0.9682	0.9462
10	1.0000	1.0000	1.0000	0.9999	0.9998	0.9997	0.9995	0.9992	0.9987	0.9981	0.9972	0.9933	0.9863	0.9747
11	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9999	0.9998	0.9996	0.9994	0.9991	0.9976	0.9945	0.9890
12	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9999	0.9998	0.9997	0.9992	0.9980	0.9955
13	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9997	0.9993	0.9983
14	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9994
15	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998
16	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
17	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000

	6.0	6.5	7.0	7.5	8.0	8.5	9.0	9.5	10.0	11.0	12.0	13.0	14.0	15.0
0	0.0025	0.0015	0.0009	0.0006	0.0003	0.0002	0.0001	0.0001	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
1	0.0174	0.0113	0.0073	0.0047	0.0030	0.0019	0.0012	0.0008	0.0005	0.0002	0.0005	0.0001	0.0000	0.0000
2	0.0620	0.0430	0.0296	0.0203	0.0138	0.0093	0.0062	0.0042	0.0028	0.0012	0.0028	0.0005	0.0001	0.0000
3	0.1512	0.1118	0.0818	0.0591	0.0424	0.0301	0.0212	0.0149	0.0103	0.0049	0.0103	0.0023	0.0005	0.0002
4	0.2851	0.2237	0.1730	0.1321	0.0996	0.0744	0.0550	0.0403	0.0293	0.0151	0.0293	0.0076	0.0018	0.0009
5	0.4457	0.3690	0.3007	0.2414	0.1912	0.1496	0.1157	0.0885	0.0671	0.0375	0.0671	0.0203	0.0055	0.0028
6	0.6063	0.5265	0.4497	0.3782	0.3134	0.2562	0.2068	0.1649	0.1301	0.0786	0.1301	0.0458	0.0142	0.0076
7	0.7440	0.6728	0.5987	0.5246	0.4530	0.3856	0.3239	0.2687	0.2202	0.1432	0.2202	0.0895	0.0316	0.0180
8	0.8472	0.7916	0.7291	0.6620	0.5925	0.5231	0.4557	0.3918	0.3328	0.2320	0.3328	0.1550	0.0621	0.0374
9	0.9161	0.8774	0.8305	0.7764	0.7166	0.6530	0.5874	0.5218	0.4579	0.3405	0.4579	0.2424	0.1094	0.0699
10	0.9574	0.9332	0.9015	0.8622	0.8159	0.7634	0.7060	0.6453	0.5830	0.4599	0.5830	0.3472	0.1757	0.1185
11	0.9799	0.9661	0.9467	0.9208	0.8881	0.8487	0.8030	0.7520	0.6968	0.5793	0.6968	0.4616	0.2600	0.1848
12	0.9912	0.9840	0.9730	0.9573	0.9362	0.9091	0.8758	0.8364	0.7916	0.6887	0.7916	0.5760	0.3585	0.2676
13	0.9964	0.9929	0.9872	0.9784	0.9658	0.9486	0.9261	0.8981	0.8645	0.7813	0.8645	0.6815	0.4644	0.3632
14	0.9986	0.9970	0.9943	0.9897	0.9827	0.9726	0.9585	0.9400	0.9165	0.8540	0.9165	0.7720	0.5704	0.4657
15	0.9995	0.9988	0.9976	0.9954	0.9918	0.9862	0.9780	0.9665	0.9513	0.9074	0.9513	0.8444	0.6694	0.5681
16	0.9998	0.9996	0.9990	0.9980	0.9963	0.9934	0.9889	0.9823	0.9730	0.9441	0.9730	0.8987	0.7559	0.6641
17	0.9999	0.9998	0.9996	0.9992	0.9984	0.9970	0.9947	0.9911	0.9857	0.9678	0.9857	0.9370	0.8272	0.7489
18	1.0000	0.9999	0.9999	0.9997	0.9993	0.9987	0.9976	0.9957	0.9928	0.9823	0.9928	0.9626	0.8826	0.8195
19	1.0000	1.0000	1.0000	0.9999	0.9997	0.9995	0.9989	0.9980	0.9965	0.9907	0.9965	0.9787	0.9235	0.8752
20	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9996	0.9991	0.9984	0.9953	0.9984	0.9884	0.9521	0.9170
21	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998	0.9996	0.9993	0.9977	0.9993	0.9939	0.9712	0.9469
22	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9999	0.9997	0.9990	0.9997	0.9970	0.9833	0.9673
23	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9999	0.9995	0.9999	0.9985	0.9907	0.9805
24	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9998	1.0000	0.9993	0.9950	0.9888
25	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	1.0000	0.9997	0.9974	0.9938
26	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9987	0.9967
27	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9994	0.9983
28	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9997	0.9991
29	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9996
30	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9998
31	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999
32	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000